# The Yule Walker Equations for the AR Coefficients

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If you assume a given zero-mean discrete timeseries  $\{x_i\}_1^N$  is an AR process, you will naturally want to estimate the appropriate order p of the AR(p),

$$x_{i+1} = \phi_1 x_i + \phi_2 x_{i-1} + \dots + \phi_p x_{i-p+1} + \xi_{i+1}$$
(1)

and the corresponding coefficients  $\{\phi_j\}$ . There are (at least) 2 methods, and those are described in this section.

## 1 Direct Inversion

The first possibility is to form a set of direct inversions,

**1.1** p = 1

With

$$x_{i+1} = \phi_1 x_i + \xi_{i+1},$$

one can form the over-determined system

$$\underbrace{\begin{pmatrix} x_2 \\ x_3 \\ \vdots \\ x_N \end{pmatrix}}_{\mathbf{b}} = \underbrace{\begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_{N-1} \end{pmatrix}}_{\mathbf{A}} \phi_1$$

which can be readily solve using the usual least-squares estimator

$$\hat{\phi}_1 = \left(\mathbf{A}^T \mathbf{A}\right)^{-1} \mathbf{A}^T \mathbf{b} = \frac{\sum_{i=1}^{N-1} x_i x_{i+1}}{\sum_{i=1}^{N-1} x_i^2} = \frac{c_1}{c_o} = r_1$$

where  $c_i$  and  $r_i$  are the ith autocovariance and autocorrelation coefficients, respectively.

**1.2** 
$$p = 2$$

With

$$x_{i+1} = \phi_1 x_i + \phi_2 x_{i-1} + \xi_{i+1},$$

start by forming the over-determined system

$$\begin{pmatrix} x_3 \\ x_4 \\ \vdots \\ x_N \end{pmatrix} = \underbrace{\begin{pmatrix} x_2 & x_1 \\ x_3 & x_2 \\ \vdots & \vdots \\ x_{N-1} & x_{N-2} \end{pmatrix}}_{\mathbf{A}} \underbrace{\begin{pmatrix} \phi_1 \\ \phi_2 \end{pmatrix}}_{\mathbf{\Phi}}.$$

Unlike the previous p = 1 case, trying to express the solution

$$\hat{\mathbf{\Phi}} = \left(\mathbf{A}^T \mathbf{A}\right)^{-1} \mathbf{A}^T \mathbf{b}$$

analytically is not trivial. We start with

$$\left( \mathbf{A}^{T} \mathbf{A} \right)^{-1} = \begin{bmatrix} \begin{pmatrix} x_{2} & x_{3} & \cdots & x_{N-1} \\ x_{1} & x_{2} & \cdots & x_{N-2} \end{pmatrix} \begin{pmatrix} x_{2} & x_{1} \\ x_{3} & x_{2} \\ x_{N-1} & x_{N-2} \end{pmatrix} \end{bmatrix}^{-1} \\ = \begin{pmatrix} \sum_{i=2}^{N-1} x_{i}^{2} & \sum_{i=2}^{N-1} x_{i} x_{i-1} \\ \sum_{i=2}^{N-1} x_{i} x_{i-1} & \sum_{i=1}^{N-2} x_{i}^{2} \end{pmatrix}^{-1} \\ = \frac{1}{\sum_{i=2}^{N-1} x_{i}^{2} \sum_{i=1}^{N-2} x_{i}^{2} - \sum_{i=2}^{N-1} x_{i} x_{i-1} \sum_{i=2}^{N-1} x_{i} x_{i-1}} \begin{pmatrix} \sum_{i=1}^{N-2} x_{i}^{2} & -\sum_{i=2}^{N-1} x_{i} x_{i-1} \\ -\sum_{i=2}^{N-1} x_{i} x_{i-1} & \sum_{i=2}^{N-1} x_{i}^{2} \end{pmatrix}.$$

Next, let's use the fact that the timeseries is stationary, so that autocovariance elements are a function of the lag only, not the exact time limits. In this case,

$$\left(\mathbf{A}^{T}\mathbf{A}\right)^{-1} = \frac{1}{c_{o}^{2} - c_{1}^{2}} \begin{pmatrix} c_{o} & -c_{1} \\ -c_{1} & c_{o} \end{pmatrix},$$
$$\left(\mathbf{A}^{T}\mathbf{A}\right)^{-1} = \frac{1}{c_{o}^{2}(1 - r_{1}^{2})} \begin{pmatrix} c_{o} & -c_{1} \\ -c_{1} & c_{o} \end{pmatrix},$$
$$\left(\mathbf{A}^{T}\mathbf{A}\right)^{-1} = \frac{1}{c_{o}(1 - r_{1}^{2})} \begin{pmatrix} r_{o} & -r_{1} \\ -r_{1} & r_{o} \end{pmatrix}.$$

Similarly,

$$\mathbf{A}^{T}\mathbf{b} = \begin{pmatrix} x_{2} & x_{3} & \cdots & x_{N-1} \\ x_{1} & x_{2} & \cdots & x_{N-2} \end{pmatrix} \begin{pmatrix} x_{3} \\ x_{4} \\ \vdots \\ x_{N} \end{pmatrix} = \begin{pmatrix} \sum_{i=3}^{N} x_{i}x_{i-1} \\ \sum_{i=3}^{N} x_{i}x_{i-2}, \end{pmatrix}$$

which, exploiting again the stationarity of the timeseries, becomes

$$\mathbf{A}^T \mathbf{b} = \left(\begin{array}{c} c_1 \\ c_2 \end{array}\right).$$

Combining the 2 expressions, we have

$$\left( \mathbf{A}^T \mathbf{A} \right)^{-1} \mathbf{A}^T \mathbf{b} = \frac{1}{c_o (1 - r_1^2)} \begin{pmatrix} r_o & -r_1 \\ -r_1 & r_o \end{pmatrix} \begin{pmatrix} c_1 \\ c_2 \end{pmatrix}$$
$$= \frac{1}{1 - r_1^2} \begin{pmatrix} 1 & -r_1 \\ -r_1 & 1 \end{pmatrix} \begin{pmatrix} r_1 \\ r_2 \end{pmatrix}.$$

Breaking this into individual components, we get

$$\hat{\phi}_1 = \frac{r_1 \left(1 - r_2\right)}{1 - r_1^2}$$

and

$$\hat{\phi_2} = \frac{r_2 - r_1^2}{1 - r_1^2}$$

Of course it is possible to continue to explore  $p \geq 3$  cases in this fashion. However, the algebra, while not fundamentally different from the p = 2 case, quickly becomes quite nightmarish. For example, for p = 3,

$$\mathbf{A}^T \mathbf{A} = \begin{pmatrix} c_o & c_1 & c_2 \\ c_1 & c_o & c_1 \\ c_2 & c_1 & c_o \end{pmatrix},$$

whose determinant, required for the inversion, is the cumbersome-looking

$$\det \left( \mathbf{A}^T \mathbf{A} \right) = c_o \left( c_o^2 - 2c_1^2 + 2\frac{c_1^2 c_2}{c_o} - c_2^2 \right) = c_o \left[ c_o^2 + 2c_1^2 \left( r_2 - 1 \right) - c_2^2 \right],$$

which, on pre-multiplying by the remainder matrix, yields very long expressions.

Fortunately, there is a better, easier way to obtain the AR coefficient for the arbitrary p, the Yule-Walker Equations.

## 2 The Yule-Walker Equations

Consider the general AR(p)

$$x_{i+1} = \phi_1 x_i + \phi_2 x_{i-1} + \dots + \phi_p x_{i-p+1} + \xi_{i+1}.$$

#### 2.1 Lag 1

• multiply both sides of the model by  $x_i$ ,

$$x_i x_{i+1} = \sum_{j=1}^{p} \left( \phi_j x_i x_{i-j+1} \right) + x_i \xi_{i+1},$$

where i and j are the time and term indices, respectively,

• take expectance,

$$\langle x_i x_{i+1} \rangle = \sum_{j=1}^p \left( \phi_j \langle x_i x_{i-j+1} \rangle \right) + \langle x_i \xi_{i+1} \rangle$$

where the  $\{\phi_j\}$ s are kept outside the expectance operator because they are deterministic, rather than statistical, quantities.

• note that  $\langle x_i \xi_{i+1} \rangle = 0$  because the shock (or random perturbation)  $\xi$  of the current time is unrelated to-and thus uncorrelated with-previous values of the process,

$$\langle x_i x_{i+1} \rangle = \sum_{j=1}^p \left( \phi_j \langle x_i x_{i-j+1} \rangle \right)$$

• divide through by (N-1), and use the evenness of the autocovariance,  $c_{-l} = c_l$ ,

$$c_1 = \sum_{j=1}^p \phi_j c_{j-1}$$

• divide through by  $c_o$ ,

$$r_1 = \sum_{j=1}^p \phi_j r_{j-1}.$$

## 2.2 Lag 2

• multiply by  $x_{i-1}$ ,

$$x_{i-1}x_{i+1} = \sum_{j=1}^{p} \left(\phi_j x_{i-1}x_{i-j+1}\right) + x_{i-1}\xi_{i+1},$$

• take expectance,

$$\langle x_{i-1}x_{i+1}\rangle = \sum_{j=1}^{p} \left(\phi_j \langle x_{i-1}x_{i-j+1}\rangle\right) + \langle x_{i-1}\xi_{i+1}\rangle$$

• eliminate the zero correlation forcing term

$$\langle x_{i-1}x_{i+1}\rangle = \sum_{j=1}^{p} \left(\phi_j \langle x_{i-1}x_{i-j+1}\rangle\right)$$

• divide through by (N-1), and use  $c_{-l} = c_l$ ,

$$c_2 = \sum_{j=1}^p \phi_j c_{j-2}$$

• divide through by  $c_o$ ,

$$r_2 = \sum_{j=1}^p \phi_j r_{j-2}.$$

## 2.3 Lag k

• multiply by  $x_{i-k-1}$ ,

$$x_{i-k+1}x_{i+1} = \sum_{j=1}^{p} \left(\phi_j x_{i-k+1} x_{i-j+1}\right) + x_{i-k+1}\xi_{i+1},$$

• take expectance,

$$\langle x_{i-k+1}x_{i+1}\rangle = \sum_{j=1}^{p} (\phi_j \langle x_{i-k+1}x_{i-j+1}\rangle) + \langle x_{i-k+1}\xi_{i+1}\rangle$$

 $\bullet\,$  eliminate the zero correlation forcing term

$$\langle x_{i-k+1}x_{i+1}\rangle = \sum_{j=1}^{p} \left(\phi_j \langle x_{i-k+1}x_{i-j+1}\rangle\right)$$

• divide through by (N-1), and use  $c_{-l} = c_l$ ,

$$c_k = \sum_{j=1}^p \phi_j c_{j-k}$$

• divide through by  $c_o$ ,

$$r_k = \sum_{j=1}^p \phi_j r_{j-k}.$$

#### 2.4 Lag p

• multiply by  $x_{i-p-1}$ ,

$$x_{i-p+1}x_{i+1} = \sum_{j=1}^{p} \left(\phi_j x_{i-p+1} x_{i-j+1}\right) + x_{i-p+1}\xi_{i+1},$$

• take expectance,

$$\langle x_{i-p+1}x_{i+1}\rangle = \sum_{j=1}^{p} \left(\phi_j \langle x_{i-p+1}x_{i-j+1}\rangle\right) + \langle x_{i-p+1}\xi_{i+1}\rangle$$

• eliminate the zero correlation forcing term

$$\langle x_{i-p+1}x_{i+1}\rangle = \sum_{j=1}^{p} \left(\phi_j \langle x_{i-p+1}x_{i-j+1}\rangle\right)$$

• divide through by (N-1), and use  $c_{-l} = c_l$ ,

$$c_p = \sum_{j=1}^p \phi_j c_{j-p}$$

• divide through by  $c_o$ ,

$$r_p = \sum_{j=1}^p \phi_j r_{j-p}.$$

#### 2.5 Putting it All Together

Rewriting all the equations together yields

which can also be written as

$$\begin{pmatrix} r_1 \\ r_2 \\ \vdots \\ r_{p-1} \\ r_p \end{pmatrix} = \begin{pmatrix} r_o & r_1 & r_2 & \cdots & r_{p-2} & r_{p-1} \\ r_1 & r_o & r_1 & \cdots & r_{p-3} & r_{p-2} \\ \vdots & & \vdots & & \vdots \\ r_{p-2} & r_{p-3} & r_{p-4} & \cdots & r_o & r_1 \\ r_{p-1} & r_{p-2} & r_{p-3} & \cdots & r_1 & r_o \end{pmatrix} \begin{pmatrix} \phi_1 \\ \phi_2 \\ \vdots \\ \phi_{p-1} \\ \phi_p \end{pmatrix}.$$

Recalling that  $r_o = 1$ , the above equation is also

$$\underbrace{\begin{pmatrix} r_1 \\ r_2 \\ \vdots \\ r_{p-1} \\ r_p \end{pmatrix}}_{\mathbf{r}} = \underbrace{\begin{pmatrix} 1 & r_1 & r_2 & \cdots & r_{p-2} & r_{p-1} \\ r_1 & 1 & r_1 & \cdots & r_{p-3} & r_{p-2} \\ \vdots & & \vdots & & \\ r_{p-2} & r_{p-3} & r_{p-4} & \cdots & 1 & r_1 \\ r_{p-1} & r_{p-2} & r_{p-3} & \cdots & r_1 & 1 \end{pmatrix}}_{\mathbf{R}} \underbrace{\begin{pmatrix} \phi_1 \\ \phi_2 \\ \vdots \\ \phi_{p-1} \\ \phi_p \end{pmatrix}}_{\mathbf{\Phi}}$$

or succinctly

$$\mathbf{R}\boldsymbol{\Phi} = \mathbf{r}.$$
 (2)

Note that this is a well-posed system (with a square coefficients matrix  $\mathbf{R}$ ), i.e., with the same number of constraints (equations,  $\mathbf{R}$ 's rows) as unknowns (the elements  $\phi_j$  of the unknown vector  $\mathbf{\Phi}$ ). Further,  $\mathbf{R}$  is full-rank and symmetric, so that invertability is guaranteed,

$$\hat{\mathbf{\Phi}} = \mathbf{R}^{-1}\mathbf{r}.$$

## 3 The Yule-Walker Equations and the Partial Autocorrelation Function

Equation 2 provides a convenient recursion for computing the pacf. The first step is to compute the acf up to a reasonable cutoff, say  $p \simeq N/4$ . Next, let  $\mathbf{r}^{(i)}$  denote Equation 2's rhs for the p = i case. Similarly, let  $\mathbf{R}^{(i)}$  denote the coefficient matrix for the same case. Then

- loop on  $i, 1 \le i \le p$ - compute  $\mathbf{R}^{(i)}$  and  $\mathbf{r}^{(i)}$ 
  - invert for  $\hat{\Phi}^{(i)}$ ,

$$\hat{\boldsymbol{\Phi}}^{(i)} = \left(\mathbf{R}^{(i)}\right)^{-1} \mathbf{r}^{(i)} = \begin{pmatrix} \hat{\phi}_1 \\ \hat{\phi}_2 \\ \vdots \\ \hat{\phi}_i \end{pmatrix}$$

- discard all  $\hat{\phi}_j$  for  $1 \le j \le i - 1$ - retain  $\hat{\phi}_i$ ,

$$\operatorname{pacf}(i) = \hat{\phi}_i$$

- $\bullet$  end loop on i
- plot pacf(i) as a function of i.