

## *Categorical Predictors and Interaction*

### Preliminaries

#### **Supplemental practice problems**

- Multiple regression practice questions

#### **Feedback and questions**

- Please continue to use the in-class forms
- Academic reps, members of cohort QC team
- E-mail and office hours

### Project

#### **Get your data downloaded this week**

- Download as text file
  - Open text file and import data into JMP-IN spreadsheet
  - Initial graphical analysis.
- Look back at analysis of first problem on Assignment #2

### Review of Key Points from Prior Class

#### **Confounding effects**

- Flaws in a marginal comparison  
When the observations (in that example, the managers) have not been randomly assigned to the groups, *confounding* is possible.
- Confounding, in general  
The differences that we detect in a marginal comparison are due to other factors, *not* the nominal labeling that defines the groups. Confounding is collinearity in another guise (e.g., it's age not shoe size in IQ example).
- Confounding, in the hiring application  
Internal managers look worse marginally, but better when adjusted to be doing comparable jobs (i.e., in the sense of jobs that pay equally).
- Why does it matter that the factors are confounded?  
Consider the effects of actions taken based on the marginal comparison. In this application, you'd hire the less desirable employee (on average).

## Adjusting for confounding

- Compare predictions of two regression models  
Which group has the higher expected performance rating when filling a specific position in the company. The fit for the *internal* managers is higher.
- Question  
Are the fitted regressions significantly different, or have we found a random difference rather than a systematic difference? Would the same effect show up in another sample, or is it just a property of this sample?
- Parallel  
The comparison of the two regressions is “easy” if the fits are parallel, for then the difference when hiring a manager at \$75,000 is the same as hiring one at \$85,000 – or at zero for that matter. Thus we can focus on the difference of the intercepts, *if the fits are parallel*.

## Common Question

### Why not compare the intercepts using their confidence intervals?

- Recall the comparison of two averages using confidence intervals rather than the two-sample t-test in Stat 603.
- Such a procedure is OK when the differences are large, but...
  - (1) This procedure does not generalize naturally to problems having more than two groups (gets pretty tedious then).
  - (2) This procedure will miss some differences that are significant.

## Key Applications for Today

### Modeling with interaction terms

- Are the fitted slopes for internal and external managers different?
- Do spending patterns differ among consumers from different nations?

## Concepts and Terminology

### Combining several regressions into one

- How do we test for differences between the regressions fit to separate groups?
- It would be easy if we had
  - (1) a t-ratio for the difference in slopes and
  - (2) another for the difference in intercepts.
- “Glue” the separate simple regressions into *one* multiple regression.

### Interaction

- Question: Does each predictor affect the response in the same way (i.e., have the same slope) for each group?
- Rephrased: Is there an *interaction* between the predictor and the categorical variable that identifies the groups.
- Consequence: Interaction complicates interpretation since the fits are no longer parallel, and thus the difference between fits is not constant.

### Important questions to answer when using categorical variables

- Are the groups really so similar as to make sense to combine the data?
  - Are the fits in the different models parallel? (i.e., Is interaction present?)
  - If they are parallel, are the intercepts different?
  - Are the error variances comparable? (i.e., Is heteroscedasticity present?)
- The multiple regression that joins the fits to the separate groups has one RMSE. Each fit to separate groups has its own. Is one good enough?

## JMP-IN

### Reading the output

- Each regression with categorical predictors has a “baseline” model common to all groups.
- Terms associated with a categorical predictor shift the equation up and down by changing the intercept.
- Interactions that combine categorical and continuous predictors shift the slope from the baseline model.

### **It looks easy in class, but when I try to do it, I can't figure out where to click next. What should I do?**

Take a look in the index of the JMP-IN manual. It has a summary of all the menu commands. Try the on-line help too.

Also, keeping up with the class is easier if you not only skim the cases, but also try to reproduce the output as well. Then you'll know what to look for during the class demonstration.

### **How do you fit separate lines to each group in a scatterplot?**

Use the red triangle pop-up menu at the upper left of the window. Select the “Group by” option and choose the categorical factor of interest. When you next fit a line in that scatterplot, JMP-IN will fit a separate line for each group.

### **How do you see results for both categories?**

Be sure to ask JMP-IN for the *expanded* estimates (using the red triangle to get to the pop-up dialog in the multiple regression output).

### **How do you color-code the points?**

Use the “Color/mark by Column” command from the *Rows* menu.

### **How do you add an interaction to a regression?**

An interaction involves a pair of predictors. Select the two that you want to combine and click the “cross” button. JMP-IN will form the interaction term as an added predictor (labeled, e.g., as Origin\*Manager). It does not matter which name comes first.

## Examples for Today

### Employee performance study

Manager.jmp, page 161

“Which of two prospective job candidate should we hire, the internal or the externally recruited manager?”

#### Data

150 managers, 88 of which are internal and 62 are external.

#### Variables

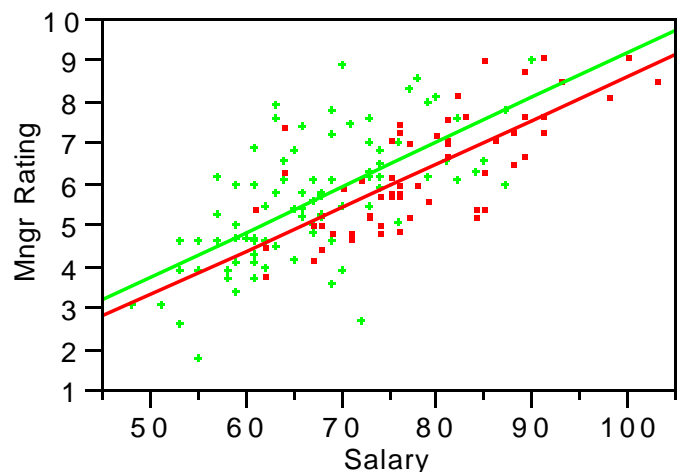
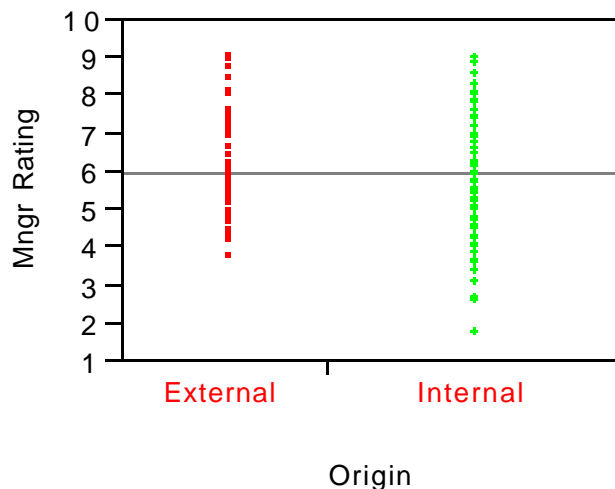
- *Salary* is the starting salary of the employee when hired and indicates what sort of job the person was hired to do.
- *MngrRating* is an evaluation score of the employee in the job they do, indicating the “value” of the employee to the firm. Higher paying jobs typically earn higher ratings since such jobs provide the opportunity to offer more important contributions to the firm.

#### Analysis

In a preliminary marginal analysis, the average performance rating for **external** managers is significantly higher than that for **internal** managers  
(avg ext.– avg int.) = 0.72 with  $t = 2.98$  (page 161)

*Confounding issue...*

*Salary* is higher for **externally** recruited managers... They occupy higher level positions within the company, and *Salary* is related to rating (p 164-165).



*Separate fits...*

Separate regressions of *Rating* on *Salary* for *In-House?* Reverse the difference: at a fixed salary, **internal** are more highly rated! (p166-67) The green line in the figure for **internal** managers is consistently higher than the red line giving the expected rating for **external** managers.

*Statistical test of difference*

If we assume slopes are parallel (i.e., no evident interaction), a model using a categorical predictor (expanded estimates version of model on page 168) implies that **internal** managers actually rate *significantly* higher, if we assume that the slopes are the same.

<b>Expanded Estimates</b>					
<b>Term</b>	<b>Estimate</b>	<b>Std Error</b>	<b>t Ratio</b>	<b>Prob&gt; t </b>	
Intercept	-1.84	0.71	-2.61	0.0100	
Salary	0.107	0.01	11.14	<.0001	
Origin[External]	-0.257	0.10	-2.46	0.0149	
Origin[Internal]	0.257	0.10	<b>2.46</b>	0.0149	

*Fits for two groups*

Internal Predicted rating =  $-1.84 + 0.107 \text{ Salary} + 0.257$

External Predicted rating =  $-1.84 + 0.107 \text{ Salary} - 0.257$

difference in intercepts =  $-0.514 = 0.257 - (-0.257)$  with  $t = -2.46$ , twice the coefficient in the regression summary.

*Are the fitted models are parallel?*

Rather than look at the picture, fit a model that allows the slopes to differ. Use it to estimate the difference between the slopes. The key is to use an *interaction*. Add it to the model using the “cross” button. (see the compact output on p. 173). In this example, the interaction is *not* significant

<b>Expanded Estimates</b>					
<b>Term</b>	<b>Estimate</b>	<b>Std Error</b>	<b>t Ratio</b>	<b>Prob&gt; t </b>	
Intercept	-1.815	0.724	-2.51	0.0132	
Salary	0.107	0.010	10.99	<.0001	
Origin[External]	-0.254	0.106	-2.39	0.0179	
Origin[Internal]	0.254	0.106	2.39	0.0179	
(Salary-71.63)*Origin[External]	-0.0018	0.010	-0.19	0.8499	
(Salary-71.63)*Origin[Internal]	0.0018	0.010	0.19	0.8499	

*Fit for internal group*

$$\begin{aligned}\text{Predicted rating} &= -1.815 + 0.107 \text{ Salary} \\ &\quad + 0.254 + 0.0018 (\text{Salary} - 71.63) \\ &= -1.815 + 0.254 - 0.0018(71.63) + (0.107 + 0.0018) \text{ Salary} \\ &= -1.690 + 0.109 \text{ Salary}\end{aligned}$$

This fit *is* the initial simple regression for this group – but now we can compare it to the fit for the external group using the t-ratios in the output.

*Note on the output*

JMP-IN forms an interaction by subtracting the mean from the continuous predictor. Doing so reduces collinearity in this type of regression model.

### **Conclude**

After checking assumptions (particularly, the assumption of equal error variance on p. 174), conclude that the company ought to hire the **internal** candidate since at any given salary, we expect the **internal** manager to rate better by a statistically significant margin.

*Comment on the order of our analysis*

This type of analysis would usually begin by first checking to see if the slopes are parallel before looking at the difference of intercepts. Starting with the full model that includes interactions is a bit hard from a teaching point of view, so I have built up to the full model rather than start there.

## Supplemental Wage discrimination

Salary.jmp, page 180

“Are men paid more than women in mid-level management positions?”

### Data

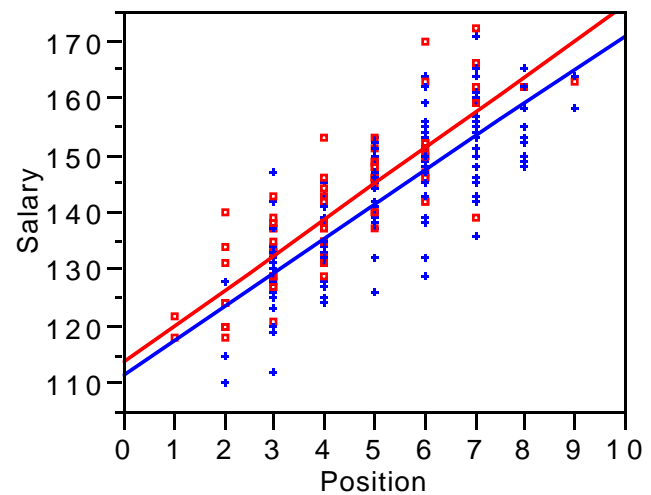
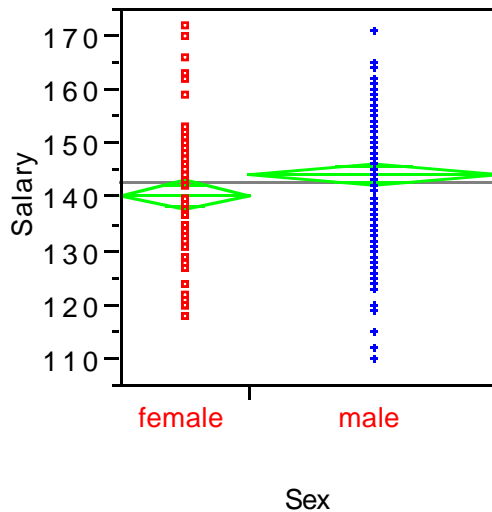
220 managers, 145 are men and 75 are women.

### Marginal analysis

Men are paid more (p 181), to the tune of about \$3600 per year. This difference is statistically significant.

### Confounding

More “careful” analysis shows that men occupy higher-level positions than women in this sample. If the analysis is restricted to a subset that conditions on the position of the manager, the revised analysis indicates that men are paid less, not more. The loss of data leads to insignificant effect. (p 183)





*Regression with categorical predictor*

Uses of all of the data rather than a subset and finds that the salary difference is significant (table below shows *expanded* estimates)

<b>Expanded Estimates</b>				
<b>Term</b>	<b>Estimate</b>	<b>Std Error</b>	<b>t Ratio</b>	<b>Prob&gt; t </b>
Intercept	112.77	1.45	77.52	<.0001
Position	6.06	0.28	21.60	<.0001
Sex[female]	1.86	0.53	3.53	0.0005
Sex[male]	-1.86	0.53	-3.53	0.0005

*Fits for the two groups*

Women: predicted salary = 112.77 + 6.06 Position + 1.86

Men: predicted salary = 112.77 + 6.06 Position - 1.86

Thus we find that a man on average *at a given position* (rather than marginally) is being paid about  $2(1.86)1000 = \$3,720$  less than a woman in the same position.

*Check for interaction*

This fit assumes that the two population lines are parallel, thus it forces the fitted lines to be parallel as well. To check this assumption, add an interaction (though it again looks clear there is no interaction).

<b>Expanded Estimates</b>				
<b>Term</b>	<b>Estimate</b>	<b>Std Error</b>	<b>t Ratio</b>	<b>Prob&gt; t </b>
Intercept	112.63	1.48	75.85	<.0001
Position	6.10	0.30	20.58	<.0001
Sex[female]	1.92	0.54	3.54	0.0005
Sex[male]	-1.92	0.54	-3.54	0.0005
Sex[female]*(Position-5.068)	0.15	0.30	0.49	0.6242
Sex[male]*(Position-5.068)	-0.15	0.30	-0.49	0.6242

*Fits for the two groups*

The model that includes both the categorical predictor and the interaction allows both the slopes and intercepts to differ (reproducing the original separate fits, but within one regression model)

$$\begin{aligned} \text{Women:} \quad & \text{predicted salary} = 112.63 + 6.10 \text{ Position} + 1.92 \\ & \quad \quad \quad + 0.15 (\text{Position} - \text{avg Position}) \\ \text{Men:} \quad & \text{predicted salary} = 112.63 + 6.10 \text{ Position} - 1.92 \\ & \quad \quad \quad - 0.15 (\text{Position} - \text{avg Position}) \end{aligned}$$

*Interaction is not significant*

The t-ratio, confidence interval, or p-value all imply that the addition of the interaction term has not improved the fit of the model, so in the interest of *parsimony*, we ought to remove it and work with the simpler model.

### **Conclude**

After checking assumptions (are variances comparable in these two groups) we find that a comparison of men with women having statistically adjusted backgrounds suggest that men are paid less. However, why do the women occupy lower level positions?

## **Key Take-Away Points**

### **Categorical predictors in regression**

- Test for differences between regression models fit to two groups.  
By combining the two groups and fitting one multiple regression rather than two simple regressions, we can test for differences between
  - the slopes (i.e., test for interaction)and between
  - the intercepts (test for shifts in the fit)
- Categorical terms alter intercept, interactions alter the associated slope.
- Relevant error assumption to check
  - Is the variation about the fit the same in all groups.

## **Next Time**

**Regression with more than two categories.**

**Building a regression model.**